# MATH 532: Linear Algebra

## Chapter 4: Vector Spaces

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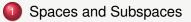
# Outline



- Pour Fundamental Subspaces
  - Linear Independence
  - Bases and Dimension
  - More About Rank
- 6 Classical Least Squares
  - Kriging as best linear unbiased predictor



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- 2) Four Fundamental Subspaces
- Linear Independence
- 4 Bases and Dimension
- 5 More About Rank
- 6 Classical Least Squares
  - Kriging as best linear unbiased predictor



# Spaces and Subspaces

While the discussion of vector spaces can be rather dry and abstract, they are an essential tool for describing the world we work in, and to understand many practically relevant consequences.

After all, linear algebra is pretty much the workhorse of modern applied mathematics.

Moreover, many concepts we discuss now for traditional "vectors" apply also to vector spaces of functions, which form the foundation of functional analysis.



# **Vector Space**

## Definition

A set  ${\cal V}$  of elements (vectors) is called a vector space (or linear space) over the scalar field  ${\cal F}$  if

(A1)  $\boldsymbol{x} + \boldsymbol{y} \in \mathcal{V}$  for any  $\boldsymbol{x}, \boldsymbol{y} \in \mathcal{V}$  (closed under addition),

(A2) 
$$(x + y) + z = x + (y + z)$$
 for all   
  $x, y, z \in \mathcal{V}$ ,

(A3) 
$$\boldsymbol{x} + \boldsymbol{y} = \boldsymbol{y} + \boldsymbol{x}$$
 for all  $\boldsymbol{x}, \boldsymbol{y} \in \mathcal{V}$ ,

- (A4) There exists a zero vector  $\mathbf{0} \in \mathcal{V}$ such that  $\mathbf{x} + \mathbf{0} = \mathbf{x}$  for every  $\mathbf{x} \in \mathcal{V}$ ,
- (A5) For every  $\boldsymbol{x} \in \mathcal{V}$  there is a negative  $(-\boldsymbol{x}) \in \mathcal{V}$  such that  $\boldsymbol{x} + (-\boldsymbol{x}) = \boldsymbol{0}$ ,

(M1)  $\alpha \mathbf{x} \in \mathcal{V}$  for every  $\alpha \in \mathcal{F}$  and  $\mathbf{x} \in \mathcal{V}$  (closed under scalar multiplication),

M2) 
$$(\alpha\beta)\mathbf{x} = \alpha(\beta\mathbf{x})$$
 for all  $\alpha\beta \in \mathcal{F}$ ,  
 $\mathbf{x} \in \mathcal{V}$ ,

(M3)  $\alpha(\boldsymbol{x} + \boldsymbol{y}) = \alpha \boldsymbol{x} + \alpha \boldsymbol{y}$  for all  $\alpha \in \mathcal{F}$ ,  $\boldsymbol{x}, \boldsymbol{y} \in \mathcal{V}$ ,

(M4) 
$$(\alpha + \beta)\mathbf{x} = \alpha \mathbf{x} + \beta \mathbf{x}$$
 for all  $\alpha, \beta \in \mathcal{F}, \mathbf{x} \in \mathcal{V},$ 

(M5) 
$$1 \mathbf{x} = \mathbf{x}$$
 for all  $\mathbf{x} \in \mathcal{V}$ .

**MATH 532** 





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But also

- $\mathcal{V}$  is polynomials of a certain degree with real coefficients,  $\mathcal{F} = \mathbb{R}$
- $\mathcal{V}$  is continuous functions on an interval  $[a, b], \mathcal{F} = \mathbb{R}$



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#### Theorem

The subset  $\mathcal{S} \subseteq \mathcal{V}$  is a subspace of  $\mathcal{V}$  if and only if

$$\alpha \mathbf{x} + \beta \mathbf{y} \in S$$
 for all  $\mathbf{x}, \mathbf{y} \in S, \ \alpha, \beta \in F$ .



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The subset  $S \subseteq V$  is a subspace of V if and only if

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#### Remark

 $\mathcal{Z} = \{\mathbf{0}\}$  is called the trivial subspace.

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If  $\boldsymbol{x} \in \mathcal{S}$ , then — using (M1) —  $-1\boldsymbol{x} = -\boldsymbol{x} \in \mathcal{S}$ , i.e., (A5) holds.



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Using (A1),  $\mathbf{x} + (-\mathbf{x}) = \mathbf{0} \in S$ , so that (A4) holds.



Let  $S = {\boldsymbol{v}_1, \dots, \boldsymbol{v}_r} \subseteq \mathcal{V}$ . The span of S is

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- If S = {v<sub>1</sub>} ⊆ ℝ<sup>3</sup>, then span(S) is the line through the origin with direction v<sub>1</sub>.
- **2** If  $S = {\mathbf{v}_1, \mathbf{v}_2 : \mathbf{v}_1 \neq \alpha \mathbf{v}_2, \alpha \neq 0} \subseteq \mathbb{R}^3$ , then span(S) is the plane through the origin "spanned by"  $\mathbf{v}_1$  and  $\mathbf{v}_2$ .

# Let $S = {\mathbf{v}_1, ..., \mathbf{v}_r} \subseteq V$ . If span S = V then S is called a spanning set for V.



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## Remark

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$$\left\{ \begin{pmatrix} 1\\0\\0 \end{pmatrix}, \begin{pmatrix} 0\\1\\0 \end{pmatrix}, \begin{pmatrix} 0\\0\\1 \end{pmatrix} \right\}$$
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$$\left\{ \begin{pmatrix} 1\\0\\0 \end{pmatrix}, \begin{pmatrix} 0\\1\\0 \end{pmatrix}, \begin{pmatrix} 0\\0\\1 \end{pmatrix} \right\} \text{ is a spanning set for } \mathbb{R}^3.$$
  
• 
$$\left\{ \begin{pmatrix} 1\\0\\0\\0 \end{pmatrix}, \begin{pmatrix} 0\\1\\0 \end{pmatrix}, \begin{pmatrix} 0\\0\\1 \end{pmatrix}, \begin{pmatrix} 2\\0\\0 \end{pmatrix}, \begin{pmatrix} 0\\2\\0 \end{pmatrix}, \begin{pmatrix} 0\\0\\2 \end{pmatrix} \right\} \text{ is also a spanning set for } \mathbb{R}^3.$$

# Connection to linear systems

#### Theorem

Let  $S = \{a_1, a_2, ..., a_n\}$  be the set of columns of an  $m \times n$  matrix A. span $(S) = \mathbb{R}^m$  if and only if for every  $\mathbf{b} \in \mathbb{R}^m$  there exists an  $\mathbf{x} \in \mathbb{R}^n$ such that  $A\mathbf{x} = \mathbf{b}$  (i.e., if and only if  $A\mathbf{x} = \mathbf{b}$  is consistent for every  $\mathbf{b} \in \mathbb{R}^m$ ).



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#### Proof.

By definition, S is a spanning set for  $\mathbb{R}^m$  if and only if for every  $\boldsymbol{b} \in \mathbb{R}^m$  there exist  $\alpha_1, \ldots, \alpha_n \in \mathbb{R}$  such that

$$m{b} = lpha_1 m{a}_1 + \ldots + lpha_n m{a}_n = m{A}m{x},$$
  
where  $m{A} = \begin{pmatrix} m{a}_1 & m{a}_2 & \cdots & m{a}_n \end{pmatrix}_{m \times n}$  and  $m{x} = \begin{pmatrix} lpha_1 \\ \vdots \\ lpha_n \end{pmatrix}$ 

#### Remark

The sum

$$\mathcal{X} + \mathcal{Y} = \{ \mathbf{x} + \mathbf{y} : \mathbf{x} \in \mathcal{X}, \mathbf{y} \in \mathcal{Y} \}$$

is a subspace of  $\mathcal{V}$  provided  $\mathcal{X}$  and  $\mathcal{Y}$  are subspaces.



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If  $S_X$  and  $S_Y$  are spanning sets for X and Y, respectively, then  $S_X \cup S_Y$  is a spanning set for X + Y.



## Outline





- Linear Independence
- 4 Bases and Dimension
- 5 More About Rank
- 6 Classical Least Squares





## Recall that a linear function $f : \mathbb{R}^n \to \mathbb{R}^m$ satisfies

 $f(\alpha \mathbf{x} + \beta \mathbf{y}) = \alpha f(\mathbf{x}) + \beta f(\mathbf{y}) \qquad \forall \alpha, \beta \in \mathbb{R}, \ \mathbf{x}, \mathbf{y} \in \mathbb{R}^{n}.$ 



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Example

Let A be a real  $m \times n$  matrix and

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The function *f* is linear since  $A(\alpha x + \beta y) = \alpha A x + \beta A y$ . Moreover, the range of *f*,

$$\mathcal{R}(f) = \{\mathbf{A}\mathbf{x}: \mathbf{x} \in \mathbb{R}^n\} \subseteq \mathbb{R}^m,$$

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is a subspace of  $\mathbb{R}^m$  since for all  $\alpha, \beta \in \mathbb{R}$  and  $\mathbf{x}, \mathbf{y} \in \mathbb{R}^n$ 

$$\alpha(\underbrace{\mathbf{A}\mathbf{x}}_{\in\mathcal{R}(f)}) + \beta(\underbrace{\mathbf{A}\mathbf{y}}_{\in\mathcal{R}(f)}) = \mathsf{A}(\alpha\mathbf{x} + \beta\mathbf{y}) \in \mathcal{R}(f).$$

#### Remark

For the situation in this example we can also use the terminology range of A (or image of A), i.e.,

$${m R}({\sf A})=\{{\sf A}{m x}:\;{m x}\in\mathbb{R}^n\}\subseteq\mathbb{R}^m$$



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$$R(\mathsf{A}) = \{\mathsf{A}oldsymbol{x}: oldsymbol{x} \in \mathbb{R}^n\} \subseteq \mathbb{R}^m$$

Similarly,

$$R(\mathsf{A}^{\mathsf{T}}) = \left\{\mathsf{A}^{\mathsf{T}} \boldsymbol{y} : \ \boldsymbol{y} \in \mathbb{R}^{m}\right\} \subseteq \mathbb{R}^{n}$$

is called the range of  $A^{T}$ .



## Column space and row space

Since

$$\mathbf{A}\mathbf{x} = \alpha_1 \mathbf{a}_1 + \ldots + \alpha_n \mathbf{a}_n,$$

we have  $R(A) = \operatorname{span}\{a_1, \dots, a_n\}$ , i.e.,

R(A) is the column space of A.



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Similarly,

 $R(A^{T})$  is the row space of A.



Consider

$$\mathsf{A} = \begin{pmatrix} 1 & 2 & 3 \\ 4 & 5 & 6 \\ 7 & 8 & 9 \end{pmatrix}$$

By definition

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In general, how do we find such minimal spanning sets as in the previous example?

An important tool is

Lemma Let A, B be  $m \times n$  matrices. Then (1)  $R(A^{T}) = R(B^{T}) \iff A \stackrel{row}{\sim} B \quad (\iff E_{A} = E_{B}).$ (2)  $R(A) = R(B) \iff A \stackrel{col}{\sim} B \quad (\iff E_{A^{T}} = E_{B^{T}}).$ 



● "⇐ ": Assume A <sup>row</sup> B, i.e., there exists a nonsingular matrix P such that

$$\mathsf{P}\mathsf{A} = \mathsf{B} \iff \mathsf{A}^T\mathsf{P}^T = \mathsf{B}^T.$$



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$$\iff \boldsymbol{a} \in \boldsymbol{R}(\boldsymbol{B}^T).$$



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2 Let 
$$A = A^T$$
 and  $B = B^T$  in (1).

#### Theorem

Let A be an  $m \times n$  matrix and U any row echelon form obtained from A. Then

- $R(A^T) = span of nonzero rows of U.$
- **2** R(A) = span of basic columns of A.



#### Theorem

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#### Remark

Later we will see that any minimal span of the columns of A forms a basis for R(A).



• This follows from (1) in the Lemma since A  $\stackrel{\text{row}}{\sim}$  U.



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- Assume the columns of A are permuted (with a matrix Q<sub>1</sub>) such that

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where B contains the basic columns, and N the nonbasic columns.

By definition, the nonbasic columns are linear combinations of the basic columns, i.e., there exists a nonsingular  $Q_2$  such that

$$\begin{bmatrix} \mathsf{B} & \mathsf{N} \end{bmatrix} \mathsf{Q}_2 = \begin{pmatrix} \mathsf{B} & \mathsf{O} \end{pmatrix},$$

where O is a zero matrix.

(cont.) Putting this together, we have

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(2) in the Lemma says that

$$R(A) = span\{B_{*1}, \ldots, B_{*r}\},\$$

where  $r = \operatorname{rank}(A)$ .



So far, we have two of the four fundamental subspaces:

R(A) and  $R(A^T)$ .



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Third fundamental subspace:  $N(A) = \{ \mathbf{x} : A\mathbf{x} = \mathbf{0} \} \subseteq \mathbb{R}^n$ ,

N(A) is the nullspace of A

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To see this, assume  $\mathbf{x}, \mathbf{y} \in N(A)$ , i.e.,  $A\mathbf{x} = A\mathbf{y} = \mathbf{0}$ . Then

$$\mathsf{A}(\alpha \boldsymbol{x} + \beta \boldsymbol{y}) = \alpha \mathsf{A} \boldsymbol{x} + \beta \mathsf{A} \boldsymbol{y} = \mathbf{0},$$

so that  $\alpha \mathbf{x} + \beta \mathbf{y} \in N(A)$ .

Find a row echelon form U of A and solve Ux = 0. Example

We can compute 
$$A = \begin{pmatrix} 1 & 2 & 3 \\ 4 & 5 & 6 \\ 7 & 8 & 9 \end{pmatrix} \longrightarrow U = \begin{pmatrix} 1 & 2 & 3 \\ 0 & -3 & -6 \\ 0 & 0 & 0 \end{pmatrix}$$
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Therefore

$$N(A) = \operatorname{span} \left\{ \begin{pmatrix} 1 \\ -2 \\ 1 \end{pmatrix} \right\}.$$

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We know 
$$rank(A) = n \iff A\mathbf{x} = \mathbf{0}$$
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**2** Repeat (1) with  $A = A^T$  and use rank( $A^T$ ) = rank(A).

## How to find a spanning set of $N(A^T)$

### Theorem

Let A be an  $m \times n$  matrix with rank(A) = r, and let P be a nonsingular matrix so that PA = U (row echelon form). Then the last m - r rows of P span  $N(A^T)$ .



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### Remark

We will later see that this spanning set is also a basis for  $N(A^T)$ .



#### Proof

# Partition P as $P = \begin{pmatrix} P_1 \\ P_2 \end{pmatrix}$ , where $P_1$ is $r \times m$ and $P_2$ is $m - r \times m$ .



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# The claim of the theorem implies that we should show that $R(P_2^T) = N(A^T)$ .



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, where  $P_1$  is  $r \times m$  and  $P_2$  is  $m - r \times m$ .

The claim of the theorem implies that we should show that  $R(P_2^T) = N(A^T)$ .

We do this in two parts:

- Show that  $R(P_2^T) \subseteq N(A^T)$ .
- **2** Show that  $N(A^T) \subseteq R(P_2^T)$ .



Partition  $U_{m \times n} = \begin{pmatrix} C \\ O \end{pmatrix}$  with  $C \in \mathbb{R}^{r \times n}$  and  $O \in \mathbb{R}^{m-r \times n}$  (a zero matrix). Then

$$\mathsf{P} \mathsf{A} = \mathsf{U} \quad \Longleftrightarrow \quad \begin{pmatrix} \mathsf{P}_1 \\ \mathsf{P}_2 \end{pmatrix} \mathsf{A} = \begin{pmatrix} \mathsf{C} \\ \mathsf{O} \end{pmatrix}$$



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This also means that

$$\mathsf{A}^{\mathsf{T}}\mathsf{P}_2^{\mathsf{T}}=\mathsf{O}^{\mathsf{T}},$$



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This also means that

$$A^T P_2^T = O^T,$$

i.e., every column of  $P_2^T$  is in  $N(A^T)$  so that  $R(P_2^T) \subseteq N(A^T)$ .



**2** Now, show  $N(A^T) \subseteq R(P_2^T)$ .



② Now, show  $N(A^T) \subseteq R(P_2^T)$ . We assume  $y \in N(A^T)$  and show that then  $y \in R(P_2^T)$ .



Now, show  $N(A^T) \subseteq R(P_2^T)$ . We assume  $y \in N(A^T)$  and show that then  $y \in R(P_2^T)$ . By definition,

$$\boldsymbol{y} \in \boldsymbol{N}(\mathsf{A}^T) \implies \mathsf{A}^T \boldsymbol{y} = \boldsymbol{0} \iff \boldsymbol{y}^T \mathsf{A} = \boldsymbol{0}^T.$$



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Since  $PA = U \implies A = P^{-1}U$ , and so

$$\mathbf{0}^{T} = \mathbf{y}^{T} \mathbf{P}^{-1} \mathbf{U} = \mathbf{y}^{T} \mathbf{P}^{-1} \begin{pmatrix} \mathbf{C} \\ \mathbf{O} \end{pmatrix}$$



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or

$$\mathbf{0}^{T} = \mathbf{y}^{T} \mathbf{Q}_{1} \mathbf{C}, \text{ where } \mathbf{P}^{-1} = \begin{pmatrix} \mathbf{Q}_{1} & \mathbf{Q}_{2} \\ m \times r & m \times m - r \end{pmatrix}$$

However, since rank(C) = r and C is  $m \times n$  we get (using m = r in our earlier theorem)

$$N(C^T) = \{\mathbf{0}\}$$

and therefore  $\boldsymbol{y}^T Q_1 = \boldsymbol{0}^T$ .



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Obviously, this implies that we also have

$$\boldsymbol{y}^T \boldsymbol{\mathsf{Q}}_1 \boldsymbol{\mathsf{P}}_1 = \boldsymbol{\mathsf{0}}^T$$



(2

(cont.) Now  $P = \begin{pmatrix} P_1 \\ P_2 \end{pmatrix}$  and  $P^{-1} = \begin{pmatrix} Q_1 & Q_2 \end{pmatrix}$  so that  $I = P^{-1}P$ 



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 $Q_1P_1=I-Q_2P_2. \label{eq:Q1}$ 



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$$\boldsymbol{y}^{T} \left( \mathsf{I} - \mathsf{Q}_{2}\mathsf{P}_{2} \right)$$



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$$\boldsymbol{y}^{T}\left(I-\boldsymbol{Q}_{2}\boldsymbol{P}_{2}\right)=\boldsymbol{0}^{T}\quad\Longleftrightarrow\quad\boldsymbol{y}^{T}=\boldsymbol{y}^{T}\boldsymbol{Q}_{2}\boldsymbol{P}_{2}$$



(cont.) Now 
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$$\boldsymbol{y}^{T}(I-Q_{2}P_{2}) = \boldsymbol{0}^{T} \quad \Longleftrightarrow \quad \boldsymbol{y}^{T} = \underbrace{\boldsymbol{y}^{T}Q_{2}}_{=\boldsymbol{z}^{T}}P_{2}.$$



(cont.) Now 
$$P = \begin{pmatrix} P_1 \\ P_2 \end{pmatrix}$$
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$$\label{eq:product} \begin{split} \boldsymbol{y}^{\mathcal{T}}\left(\mathsf{I}-\mathsf{Q}_{2}\mathsf{P}_{2}\right) = \boldsymbol{0}^{\mathcal{T}} & \Longleftrightarrow \quad \boldsymbol{y}^{\mathcal{T}} = \underbrace{\boldsymbol{y}^{\mathcal{T}}\mathsf{Q}_{2}}_{=\boldsymbol{z}^{\mathcal{T}}}\mathsf{P}_{2}. \end{split}$$
 Therefore  $\boldsymbol{y} \in R(\mathsf{P}_{2}^{\mathcal{T}}).$ 

## Finally,

# Theorem

Let A, B be  $m \times n$  matrices.

Proof. See [Mey00, Section 4.2].



# Outline



- 2 Four Fundamental Subspaces
- 3 Linear Independence
  - Bases and Dimension
  - 5 More About Rank
- 6 Classical Least Squares
  - Kriging as best linear unbiased predictor



# Linear Independence

### Definition

A set of vectors  $S = \{v_1, \dots, v_n\}$  is called linearly independent if

$$\alpha_1 \mathbf{v}_1 + \alpha_2 \mathbf{v}_2 + \ldots + \alpha_n \mathbf{v}_n = \mathbf{0} \implies \alpha_1 = \alpha_2 = \ldots = \alpha_n = \mathbf{0}.$$

Otherwise S is linearly dependent.

### Remark

Linear independence is a property of a set, not of vectors.



Is 
$$S = \left\{ \begin{pmatrix} 1 \\ 4 \\ 7 \end{pmatrix}, \begin{pmatrix} 2 \\ 5 \\ 8 \end{pmatrix}, \begin{pmatrix} 3 \\ 6 \\ 9 \end{pmatrix} \right\}$$
 linearly independent?



Is 
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Consider

$$\alpha_1 \begin{pmatrix} 1\\4\\7 \end{pmatrix} + \alpha_2 \begin{pmatrix} 2\\5\\8 \end{pmatrix} + \alpha_3 \begin{pmatrix} 3\\6\\9 \end{pmatrix} = \begin{pmatrix} 0\\0\\0 \end{pmatrix}$$



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$$\iff \mathbf{A}\mathbf{x} = \mathbf{0}, \quad \text{where } \mathbf{A} = \begin{pmatrix} 1 & 2 & 3\\4 & 5 & 6\\7 & 8 & 9 \end{pmatrix}, \ \mathbf{x} = \begin{pmatrix} \alpha_1\\\alpha_2\\\alpha_3 \end{pmatrix}$$



# Example ((cont.))

Since

$${\sf A} \stackrel{\text{row}}{\sim} {\sf E}_{{\sf A}} = egin{pmatrix} 1 & 2 & 3 \\ 0 & 1 & 2 \\ 0 & 0 & 0 \end{pmatrix}$$

we know that N(A) is nontrivial, i.e., the system Ax = 0 has a nonzero solution, and therefore S is linearly dependent.



More generally,

### Theorem

Let A be an  $m \times n$  matrix.

- The columns of A are linearly independent if and only if  $N(A) = \{0\} \iff \operatorname{rank}(A) = n.$
- **2** The rows of A are linearly independent if and only if  $N(A^T) = \{\mathbf{0}\} \iff \operatorname{rank}(A) = m.$

### Proof.

See [Mey00, Section 4.3].



### Definition

A square matrix A is called diagonally dominant if

$$|\mathbf{a}_{ii}| > \sum_{\substack{j=1\\j\neq i}}^{n} |\mathbf{a}_{ij}|, \qquad i=1,\ldots,n.$$



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### Remark

- Aside from being nonsingular (see next slide), diagonally dominant matrices are important since they ensure that Gaussian elimination will succeed without pivoting.
- Also, diagonally dominance ensures convergence of certain iterative solvers (more later).



Let A be an  $n \times n$  matrix. If A is diagonally dominant then A is nonsingular.



Let A be an  $n \times n$  matrix. If A is diagonally dominant then A is nonsingular.

### Proof

We will show that  $N(A) = \{0\}$  since then we know that rank(A) = n and A is nonsingular.

We will do this with a proof by contradiction.



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### Proof

We will show that  $N(A) = \{0\}$  since then we know that rank(A) = n and A is nonsingular.

We will do this with a proof by contradiction.

We assume that there exists an  $\mathbf{x} \neq \mathbf{0} \in N(A)$  and we will conclude that A cannot be diagonally dominant.



# (cont.) If $\boldsymbol{x} \in N(A)$ then $A\boldsymbol{x} = \boldsymbol{0}$ .



If  $\boldsymbol{x} \in N(A)$  then  $A\boldsymbol{x} = \boldsymbol{0}$ .

Now we take k so that  $x_k$  is the maximum (in absolute value) component of x and consider

 $A_{k*} x = 0.$ 



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We can rewrite this as

$$\sum_{j=1}^n a_{kj} x_j = 0 \quad \Longleftrightarrow$$



If  $\boldsymbol{x} \in N(A)$  then  $A\boldsymbol{x} = \boldsymbol{0}$ .

Now we take k so that  $x_k$  is the maximum (in absolute value) component of x and consider

$$A_{k*}\boldsymbol{x}=0.$$

We can rewrite this as

$$\sum_{j=1}^n a_{kj} x_j = 0 \quad \Longleftrightarrow \quad a_{kk} x_k = -\sum_{j=1 \atop j \neq k}^n a_{kj} x_j.$$



Now we take absolute values:

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Finally, dividing both sides by  $|x_k|$  yields

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which shows that A cannot be diagonally dominant (which is a contradiction since A was assumed to be diagonally dominant).

Consider *m* real numbers  $x_1, \ldots, x_m$  such that  $x_i \neq x_j$ ,  $i \neq j$ . Show that the columns of the Vandermonde matrix

$$\mathsf{V} = \begin{pmatrix} 1 & x_1 & x_1^2 & \cdots & x_1^{n-1} \\ 1 & x_2 & x_2^2 & \cdots & x_2^{n-1} \\ & & \vdots & \\ 1 & x_m & x_m^2 & \cdots & x_m^{n-1} \end{pmatrix}$$

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form a linearly independent set provided  $n \le m$ .

From above, the columns of V are linearly independent if and only if  $N(V) = \{\mathbf{0}\}$ 

$$\iff \forall \boldsymbol{z} = \boldsymbol{0} \implies \boldsymbol{z} = \boldsymbol{0}, \quad \boldsymbol{z} = \begin{pmatrix} \alpha_0 \\ \vdots \\ \alpha_{n-1} \end{pmatrix}$$

(cont.) Now  $V \boldsymbol{z} = \boldsymbol{0}$  if and only if

$$\alpha_0 + \alpha_1 x_i + \alpha_2 x_i^2 + \ldots + \alpha_{n-1} x_i^{n-1} = 0, \quad i = 1, \ldots, m.$$

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In other words,  $x_1, x_2, \ldots, x_m$  are all (distinct) roots of

$$p(x) = \alpha_0 + \alpha_1 x + \alpha_2 x^2 + \ldots + \alpha_{n-1} x^{n-1}$$

This is a polynomial of degree at most n - 1.

It can have *m* distinct roots only if  $m \le n - 1$ .

Otherwise, *p* is the zero polynomial, i.e.,  $\alpha_0 = \alpha_1 = \ldots = \alpha_{n-1} = 0$ , so that the columns of V are linearly dependent.

The example implies that in the special case m = n there is a unique polynomial of degree (at most) m - 1 that interpolates the data  $\{(x_1, y_1), (x_2, y_2), \dots, (x_m, y_m)\} \subset \mathbb{R}^2$ .



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$$\ell(t) = \alpha_0 + \alpha_1 t + \alpha_2 t^2 + \ldots + \alpha_{m-1} t^{m-1}.$$

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### Linear Independence

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or

$$\begin{pmatrix} 1 & x_1 & x_1^2 & \cdots & x_1^{m-1} \\ 1 & x_2 & x_2^2 & \cdots & x_2^{m-1} \\ & & \vdots & \\ 1 & x_m & x_m^2 & \cdots & x_m^{m-1} \end{pmatrix} \begin{pmatrix} \alpha_0 \\ \alpha_1 \\ \vdots \\ \alpha_{m-1} \end{pmatrix} = \begin{pmatrix} y_1 \\ y_2 \\ \vdots \\ y_m \end{pmatrix}$$

Since the columns of V are linearly independent it is nonsingular, and the coefficients  $\alpha_0, \ldots, \alpha_{m-1}$  are uniquely determined.

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In fact,

$$\ell(t) = \sum_{i=1}^{m} y_i L_i(t) \qquad \text{(Lagrange interpolation polynomial)}$$
  
with  $L_i(t) = \prod_{\substack{k=1 \ k \neq i}}^{m} (t - x_k) / \prod_{\substack{k=1 \ k \neq i}}^{m} (x_i - x_k) \qquad \text{(Lagrange functions).}$ 



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To verify (4) we note that the degree of  $\ell$  is m - 1 (since each  $L_i$  is of degree m - 1) and

$$L_i(x_j) = \delta_{ij}, \quad i, j = 1, \ldots, m,$$

so that

$$\ell(x_j) = \sum_{i=1}^m y_i \underbrace{L_i(x_j)}_{=\delta_{ij}} = y_j, \quad j = 1, \ldots, m.$$



Let  $S = \{u_1, u_2..., u_n\} \subseteq V$  be nonempty. Then

- If S contains a linearly dependent subset, then S is linearly dependent.
- If S is linearly independent, then every subset of S is also linearly independent.
- If S is linearly independent and if v ∈ V, then S<sub>ext</sub> = S ∪ {v} is linearly independent if and only if v ∉ span(S).
- If  $S \subseteq \mathbb{R}^m$  and n > m, then S must be linearly dependent.



### Proof

If S contains a linearly dependent subset, {u<sub>1</sub>,..., u<sub>k</sub>} say, then there exist nontrivial coefficients α<sub>1</sub>,..., α<sub>k</sub> such that

$$\alpha_1 \boldsymbol{u}_1 + \ldots + \alpha_k \boldsymbol{u}_k = \boldsymbol{0}.$$

Clearly, then

$$\alpha_1 \boldsymbol{u}_1 + \ldots + \alpha_k \boldsymbol{u}_k + 0 \boldsymbol{u}_{k+1} + \ldots + 0 \boldsymbol{u}_n = \boldsymbol{0}$$

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and  $\mathcal{S}$  is also linearly dependent.

Pollows from (1) by contraposition.



• " $\implies$ ": Assume  $S_{ext}$  is linearly independent. Then v can't be a linear combination of  $u_1, \ldots, u_n$ .



(cont.)

• " $\Rightarrow$ ": Assume  $S_{ext}$  is linearly independent. Then v can't be a linear combination of  $u_1, \ldots, u_n$ .

" $\Leftarrow$ ": Assume  $\boldsymbol{v} \notin \text{span}(\mathcal{S})$  and consider

 $\alpha_1 \boldsymbol{u}_1 + \alpha_2 \boldsymbol{u}_2 + \ldots + \alpha_n \boldsymbol{u}_n + \alpha_{n+1} \boldsymbol{v} = \boldsymbol{0}.$ 

First,  $\alpha_{n+1} = 0$  since otherwise  $\mathbf{v} \in \text{span}(\mathcal{S})$ .



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That leaves

$$\alpha_1 \boldsymbol{u}_1 + \alpha_2 \boldsymbol{u}_2 + \ldots + \alpha_n \boldsymbol{u}_n = \boldsymbol{0}.$$

However, the linear independence of S implies  $\alpha_i = 0$ , i = 1, ..., n, and therefore  $S_{\text{ext}}$  is linearly independent.



### (cont.)

We know that the columns of an  $m \times n$  matrix A are linearly independent if and only if rank(A) = n.

Here 
$$A = \begin{pmatrix} u_1 & u_2 & \cdots & u_n \end{pmatrix}$$
 with  $u_i \in \mathbb{R}^m$ .

If n > m, then rank(A)  $\leq m$  and S must be linearly dependent.



## Outline



- 2) Four Fundamental Subspaces
- Linear Independence
- Bases and Dimension
- More About Rank
- 6 Classical Least Squares
  - Kriging as best linear unbiased predictor



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#### Definition

Consider a vector space  $\mathcal{V}$  with spanning set  $\mathcal{S}$ . If  $\mathcal{S}$  is also linearly independent then we call it a basis of  $\mathcal{V}$ .



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- $\{e_1, \ldots, e_n\}$  is the standard basis for  $\mathbb{R}^n$ .
- Our Content of the columns/rows of an n × n matrix A with rank(A) = n form a basis for ℝ<sup>n</sup>.



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Functional analysis can be considered as infinite-dimensional linear algebra, where the linear spaces are usually function spaces such as

infinitely differentiable functions with Taylor (polynomial) basis

 $\{1, x, x^2, x^3, \ldots\}$ 

square integrable functions with Fourier basis

 $\{1, \sin(x), \cos(x), \sin(2x), \cos(2x), \ldots\}$ 



## Earlier we mentioned the idea of minimal spanning sets.



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#### Theorem

Let  $\mathcal{V}$  be a subspace of  $\mathbb{R}^m$  and let

$$\mathcal{B} = \{\boldsymbol{b}_1, \boldsymbol{b}_2, \dots, \boldsymbol{b}_n\} \subseteq \mathcal{V}.$$

The following are equivalent:

- B is a basis for V.
- **2**  $\mathcal{B}$  is a minimal spanning set for  $\mathcal{V}$ .
- $\mathbf{0}$   $\mathcal{B}$  is a maximal linearly independent subset of  $\mathcal{V}$ .



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#### Remark

We say "a basis" here since  $\mathcal{V}$  can have many different bases.



#### Proof

Since it is difficult to directly relate (2) and (3), our strategy will be to prove

- Show (1)  $\Longrightarrow$  (2) and (2)  $\Longrightarrow$  (1), so that (1)  $\iff$  (2).
- Show (1)  $\Longrightarrow$  (3) and (3)  $\Longrightarrow$  (1), so that (1)  $\iff$  (3).

Then — by transitivity — we will also have  $(2) \iff (3)$ .



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or

$$B = XA$$
,

where

$$\begin{aligned} \mathsf{B} &= \begin{pmatrix} \boldsymbol{b}_1 & \boldsymbol{b}_2 & \cdots & \boldsymbol{b}_n \end{pmatrix} \in \mathbb{R}^{m \times n}, \\ \mathsf{X} &= \begin{pmatrix} \boldsymbol{x}_1 & \boldsymbol{x}_2 & \cdots & \boldsymbol{x}_k \end{pmatrix} \in \mathbb{R}^{m \times k}, \\ [\mathsf{A}]_{ij} &= \alpha_{ij}, \quad \mathsf{A} \in \mathbb{R}^{k \times n}. \end{aligned}$$

## Proof (cont.) Now, $rank(A) \le k < n$ , which implies N(A) is



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However, since  $\mathcal{B}$  is a basis, the columns of B are linearly independent (i.e.,  $N(B) = \{0\})$  — and that is a contradiction.



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Therefore,  $\mathcal{B}$  has to be minimal.

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#### This is clear since

- if B were linearly dependent,
- then we would be able to remove at least one vector from B and still have a spanning set
- but then it would not have been minimal.



(3)  $\implies$  (1): Assume  $\mathcal{B}$  is a maximal linearly independent subset of  $\mathcal{V}$  and show that  $\mathcal{B}$  is a basis of  $\mathcal{V}$ .



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Then — by an earlier theorem — the extension set  $\mathcal{B} \cup \{v\}$  is linearly independent.



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But this contradicts the maximality of  $\mathcal{B}$ , so that  $\mathcal{B}$  has to be a basis.



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Let

$$\mathcal{Y} = \{ \boldsymbol{y}_1, \dots, \boldsymbol{y}_k \} \subseteq \mathcal{V}, \text{ with } k > n$$

be a maximal linearly independent subset of  $\mathcal{V}$  (note that such a set always exists).



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But then  $\mathcal{Y}$  must be a basis for  $\mathcal{V}$  by our "(1)  $\Longrightarrow$  (3)" argument.

On the other hand,  $\mathcal{Y}$  has more vectors than  $\mathcal{B}$  and a basis has to be a minimal spanning set.

Therefore  ${\mathcal B}$  has to already be a maximal linearly independent subset of  ${\mathcal V}.\ \ \Box$ 



Above we remarked that  $\mathcal{B}$  is not unique, i.e., a vector space  $\mathcal{V}$  can have many different bases.



#### Remark

Above we remarked that  $\mathcal{B}$  is not unique, i.e., a vector space  $\mathcal{V}$  can have many different bases.

However, the number of elements in all of these bases is unique.

# Definition The dimension of the vector space $\mathcal{V}$ is given by

 $\text{dim}\,\mathcal{V}=\text{the number of elements in any basis of}\,\,\mathcal{V}.$ 

Special case: by convention

$$\dim\{\mathbf{0}\}=0.$$

Consider

$$\mathcal{P} = \text{span} \left\{ \begin{pmatrix} 1 \\ 0 \\ 0 \end{pmatrix}, \begin{pmatrix} 0 \\ 1 \\ 0 \end{pmatrix} 
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Geometrically, P corresponds to the plane z = 0, i.e., the *xy*-plane.

Note that dim  $\mathcal{P} = 2$ .



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Geometrically,  $\mathcal{P}$  corresponds to the plane z = 0, i.e., the *xy*-plane.

Note that dim  $\mathcal{P} = 2$ .

Moreover, any subspace of  $\mathbb{R}^3$  has dimension at most 3.



#### In general,

#### Theorem

Let  $\mathcal{M}$  and  $\mathcal{N}$  be vector spaces such that  $\mathcal{M} \subseteq \mathcal{N}$ . Then

$$\mathbf{0} \quad \mathsf{dim} \, \mathcal{M} \leq \mathsf{dim} \, \mathcal{N},$$

$$e lim \mathcal{M} = \dim \mathcal{N} \implies \mathcal{M} = \mathcal{N}.$$



#### In general,

#### Theorem

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$$0 dim \mathcal{M} \leq \dim \mathcal{N},$$

$$2 \dim \mathcal{M} = \dim \mathcal{N} \implies \mathcal{M} = \mathcal{N}.$$

Proof. See [Mey00].



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 $R(A) = \text{span}\{\text{columns of } A\}.$ 

If rank(A) = r, then only r columns of A are linearly independent, i.e.,

 $\dim R(A) = r.$ 



Consider an  $m \times n$  matrix A with rank(A) = r.

R(A) We know that

 $R(A) = span\{columns of A\}.$ 

If rank(A) = r, then only r columns of A are linearly independent, i.e.,

 $\dim R(A) = r.$ 

A basis of R(A) is given by the basic columns of A (determined via a row echelon form U).



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A basis of  $R(A^T)$  is given by the nonzero rows of U (from the LU factorization of A).



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### N(A) Replace A by $A^T$ above so that

$$\dim N\left((\mathsf{A}^{\mathsf{T}})^{\mathsf{T}}\right) = n - \operatorname{rank}(\mathsf{A}^{\mathsf{T}}) = n - r$$

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A basis of N(A) is given by the n - r linearly independent solutions of  $A\mathbf{x} = \mathbf{0}$ .



#### Theorem

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This follows directly from the above discussion of R(A) and N(A).

The theorem shows that there is always a balance between the rank of A and the dimension of its nullspace.



Find the dimension and a basis for

$$\mathcal{S} = \operatorname{span} \left\{ \begin{pmatrix} 1\\2\\3\\1 \end{pmatrix}, \begin{pmatrix} 2\\4\\6\\2 \end{pmatrix}, \begin{pmatrix} 2\\4\\6\\4 \end{pmatrix}, \begin{pmatrix} 3\\6\\9\\5 \end{pmatrix}, \begin{pmatrix} 1\\2\\6\\3 \end{pmatrix} \right\}$$

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Before we even do any calculations we know that

 $\mathcal{S}\subseteq \mathbb{R}^4, \quad \text{so that } \dim \mathcal{S} \leq 4.$ 

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Before we even do any calculations we know that

$$\mathcal{S} \subseteq \mathbb{R}^4$$
, so that dim  $\mathcal{S} \leq 4$ .

We will now answer this question in two different ways using

$$\mathsf{A} = \begin{pmatrix} 1 & 2 & 2 & 3 & 1 \\ 2 & 4 & 4 & 6 & 2 \\ 3 & 6 & 6 & 9 & 6 \\ 1 & 2 & 4 & 5 & 3 \end{pmatrix}$$

.

Via R(A), i.e., by finding the basic columns of A:

$$\mathsf{A} = \begin{pmatrix} 1 & 2 & 2 & 3 & 1 \\ 2 & 4 & 4 & 6 & 2 \\ 3 & 6 & 6 & 9 & 6 \\ 1 & 2 & 4 & 5 & 3 \end{pmatrix} \quad \stackrel{\text{G.-J.}}{\longrightarrow} \quad \mathsf{E}_\mathsf{A} = \begin{pmatrix} 1 & 2 & 0 & 1 & 0 \\ 0 & 0 & 1 & 1 & 0 \\ 0 & 0 & 0 & 0 & 1 \\ 0 & 0 & 0 & 0 & 0 \end{pmatrix}$$



Via R(A), i.e., by finding the basic columns of A:

$$A = \begin{pmatrix} 1 & 2 & 2 & 3 & 1 \\ 2 & 4 & 4 & 6 & 2 \\ 3 & 6 & 6 & 9 & 6 \\ 1 & 2 & 4 & 5 & 3 \end{pmatrix} \quad \stackrel{G.-J.}{\longrightarrow} \quad E_A = \begin{pmatrix} 1 & 2 & 0 & 1 & 0 \\ 0 & 0 & 1 & 1 & 0 \\ 0 & 0 & 0 & 0 & 1 \\ 0 & 0 & 0 & 0 & 0 \end{pmatrix}$$

Therefore, dim  $\mathcal{S} = 3$  and

$$S = \operatorname{span} \left\{ \begin{pmatrix} 1 \\ 2 \\ 3 \\ 1 \end{pmatrix}, \begin{pmatrix} 2 \\ 4 \\ 6 \\ 4 \end{pmatrix}, \begin{pmatrix} 1 \\ 2 \\ 6 \\ 3 \end{pmatrix} \right\}$$

since the basic columns of  $E_A$  are the first, third and fifth columns.

Via  $R(A^{T})$ , i.e.,  $R(A) = \text{span}\{\text{rows of } A^{T}\}$ , i.e., we need the nonzero rows of U (from the LU factorization of  $A^{T}$ :

$$A^{T} = \begin{pmatrix} 1 & 2 & 3 & 1 \\ 2 & 4 & 6 & 2 \\ 2 & 4 & 6 & 4 \\ 3 & 6 & 9 & 4 \\ 1 & 2 & 6 & 3 \end{pmatrix} \xrightarrow{\text{zero out } [A^{T}]_{*,1}} \begin{pmatrix} 1 & 2 & 3 & 1 \\ 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 2 \\ 0 & 0 & 3 & 2 \end{pmatrix} \xrightarrow{\text{permute}} \underbrace{\begin{pmatrix} 1 & 2 & 3 & 1 \\ 0 & 0 & 3 & 2 \\ 0 & 0 & 0 & 2 \\ 0 & 0 & 0 & 2 \\ 0 & 0 & 0 & 2 \\ 0 & 0 & 0 & 2 \\ 0 & 0 & 0 & 2 \\ 0 & 0 & 0 & 2 \\ 0 & 0 & 0 & 2 \\ 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \\ \end{pmatrix}$$

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$$\mathcal{S} = \text{span} \left\{ \begin{pmatrix} 1 \\ 2 \\ 3 \\ 1 \end{pmatrix}, \begin{pmatrix} 0 \\ 0 \\ 3 \\ 2 \end{pmatrix}, \begin{pmatrix} 0 \\ 0 \\ 0 \\ 2 \end{pmatrix} \right\}$$

#### since the nonzero rows of U are the first, second and third rows.

=U

Example Extend

$$\mathcal{S} = \text{span} \left\{ \begin{pmatrix} 1 \\ 2 \\ 3 \\ 1 \end{pmatrix}, \begin{pmatrix} 1 \\ 2 \\ 6 \\ 3 \end{pmatrix} \right\}$$

to a basis for  $\mathbb{R}^4$ .

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The procedure will be to augment the columns of  ${\mathcal S}$  by an identity matrix , i.e., to form

$$\mathsf{A} = \begin{pmatrix} 1 & 1 & 1 & 0 & 0 & 0 \\ 2 & 2 & 0 & 1 & 0 & 0 \\ 3 & 6 & 0 & 0 & 1 & 0 \\ 1 & 3 & 0 & 0 & 0 & 1 \end{pmatrix}$$

and then to get a basis via the basic columns of U.

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$$\mathsf{A} = \begin{pmatrix} 1 & 1 & 1 & 0 & 0 & 0 \\ 2 & 2 & 0 & 1 & 0 & 0 \\ 3 & 6 & 0 & 0 & 1 & 0 \\ 1 & 3 & 0 & 0 & 0 & 1 \end{pmatrix} \ \longrightarrow$$

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$$A = \begin{pmatrix} 1 & 1 & 1 & 0 & 0 & 0 \\ 2 & 2 & 0 & 1 & 0 & 0 \\ 3 & 6 & 0 & 0 & 1 & 0 \\ 1 & 3 & 0 & 0 & 0 & 1 \end{pmatrix} \longrightarrow \begin{pmatrix} 1 & 1 & 1 & 0 & 0 & 0 \\ 0 & 0 & -2 & 1 & 0 & 0 \\ 0 & 3 & -3 & 0 & 1 & 0 \\ 0 & 2 & -1 & 0 & 0 & 1 \\ 0 & 2 & -1 & 0 & 0 & 1 \\ 0 & 0 & -\frac{3}{2} & 0 & 1 & -\frac{3}{2} \\ 0 & 0 & -2 & 1 & 0 & 0 \end{pmatrix}$$

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so that the basic columns are  $[A]_{*1}, [A]_{*2}, [A]_{*3}, [A]_{*4}$  and

$$\mathbb{R}^{4} = \text{span} \left\{ \begin{pmatrix} 1 \\ 2 \\ 3 \\ 1 \end{pmatrix}, \begin{pmatrix} 1 \\ 2 \\ 6 \\ 3 \end{pmatrix}, \begin{pmatrix} 1 \\ 0 \\ 0 \\ 0 \end{pmatrix}, \begin{pmatrix} 0 \\ 1 \\ 0 \\ 0 \end{pmatrix} \right\}$$

Earlier we defined the sum of subspaces  $\mathcal{X}$  and  $\mathcal{Y}$  as

$$\mathcal{X} + \mathcal{Y} = \{ \mathbf{x} + \mathbf{y} : \mathbf{x} \in \mathcal{X}, \mathbf{y} \in \mathcal{Y} \}$$



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#### Theorem

If  $\mathcal{X}, \mathcal{Y}$  are subspaces of  $\mathcal{V}$ , then

$$\dim(\mathcal{X} + \mathcal{Y}) = \dim \mathcal{X} + \dim \mathcal{Y} - \dim(\mathcal{X} \cap \mathcal{Y}).$$



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## Proof.

See [Mey00], but the basic idea is pretty clear. We want to avoid double counting.



# Corollary Let A and B be $m \times n$ matrices. Then

 $rank(A + B) \le rank(A) + rank(B).$ 





## Corollary

# Let A and B be $m \times n$ matrices. Then

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First we note that

$$R(A + B) \subseteq R(A) + R(B)$$

(4)

# Corollary

Let A and B be  $m \times n$  matrices. Then

$$rank(A + B) \le rank(A) + rank(B)$$
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# Proof

First we note that

$$R(A + B) \subseteq R(A) + R(B)$$

since for any  $\boldsymbol{b} \in \boldsymbol{R}(A+B)$  we have

$$\boldsymbol{b} = (\mathsf{A} + \mathsf{B})\boldsymbol{x} = \mathsf{A}\boldsymbol{x} + \mathsf{B}\boldsymbol{x} \in R(\mathsf{A}) + R(\mathsf{B}).$$



(4)

# rank(A + B) = dim R(A + B)



$$\mathsf{rank}(\mathsf{A} + \mathsf{B}) = \mathsf{dim}\,\mathsf{R}(\mathsf{A} + \mathsf{B})$$
  
 $\stackrel{(4)}{\leq} \mathsf{dim}(\mathsf{R}(\mathsf{A}) + \mathsf{R}(\mathsf{B}))$ 



$$rank(A + B) = \dim R(A + B)$$

$$\stackrel{(4)}{\leq} \dim(R(A) + R(B))$$

$$\stackrel{\text{Thm}}{=} \dim R(A) + \dim R(B) - \dim (R(A) \cap R(B))$$



```
rank(A + B) = \dim R(A + B)
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\leq \dim R(A) + \dim R(B)
= rank(A) + rank(B)
```



# Outline



- 2 Four Fundamental Subspaces
- Linear Independence
- 4 Bases and Dimension
- More About Rank
- 6 Classical Least Squares
  - Kriging as best linear unbiased predictor



We know that  $A \sim B$  if and only if rank(A) = rank(B).



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As we now show, it is a general fact that multiplication by a nonsingular matrix does not change the rank of a given matrix.



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Thus (for invertible P, Q), PAQ = B implies rank(A) = rank(PAQ).

As we now show, it is a general fact that multiplication by a nonsingular matrix does not change the rank of a given matrix.

Moreover, multiplication by an arbitrary matrix can only lower the rank.

## Theorem

Let A be an  $m \times n$  matrix, and let B by  $n \times p$ . Then

 $\operatorname{rank}(AB) = \operatorname{rank}(B) - \dim(N(A) \cap R(B)).$ 

# Remark

Note that if A is nonsingular, then  $N(A) = \{0\}$  so that dim  $(N(A) \cap R(B)) = 0$  and rank(AB) = rank(B).

#### Proof

# Let $S = \{x_1, x_2, \dots, x_s\}$ be a basis for $N(A) \cap R(B)$ .



#### Proof

Let  $S = {x_1, x_2, \dots, x_s}$  be a basis for  $N(A) \cap R(B)$ .

Since  $N(A) \cap R(B) \subseteq R(B)$  we know that

 $\dim(R(B)) = s + t$ , for some  $t \ge 0$ .



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We can construct an extension set such that

$$\mathcal{B} = \{\boldsymbol{x}_1, \boldsymbol{x}_2, \dots, \boldsymbol{x}_s, \boldsymbol{z}_1, \dots, \boldsymbol{z}_2, \dots, \boldsymbol{z}_t\}$$

is a basis for R(B).



If we can show that  $\dim(R(AB)) = t$  then

 $rank(B) = dim(R(B)) = s + t = dim(N(A) \cap R(B)) + dim(R(AB)),$ 

and we are done.

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Therefore, we now show that  $\dim(R(AB)) = t$ . In particular, we show that

$$\mathcal{T} = \{A\boldsymbol{z}_1, A\boldsymbol{z}_2, \dots, A\boldsymbol{z}_t\}$$

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$$\mathcal{T} = \{A\boldsymbol{z}_1, A\boldsymbol{z}_2, \dots, A\boldsymbol{z}_t\}$$

is a basis for R(AB).

We do this by showing that

- T is a spanning set for R(AB),
- **2** T is linearly independent.

Spanning set: Consider an arbitrary  $\boldsymbol{b} \in R(AB)$ . It can be written as

 $\boldsymbol{b} = AB\boldsymbol{y}$  for some  $\boldsymbol{y}$ .

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But then  $B\mathbf{y} \in R(B)$ , so that

$$\mathsf{B}\boldsymbol{y} = \sum_{i=1}^{s} \xi_i \boldsymbol{x}_i + \sum_{j=1}^{t} \eta_j \boldsymbol{z}_j$$

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since  $\mathbf{x}_i \in N(A)$ .

Linear independence: Let's use the definition of linear independence and look at

$$\sum_{i=1}^t \alpha_i \mathbf{A} \mathbf{z}_i = \mathbf{0} \quad \Longleftrightarrow \quad \mathbf{A} \sum_{i=1}^t \alpha_i \mathbf{z}_i = \mathbf{0}.$$

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The identity on the right implies that  $\sum_{i=1}^{t} \alpha_i \mathbf{z}_i \in N(A)$ . But we also have  $\mathbf{z}_i \in \mathcal{B}$ , i.e.,  $\sum_{i=1}^{t} \alpha_i \mathbf{z}_i \in R(B)$ .

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And so together

$$\sum_{i=1}^t \alpha_i \mathbf{z}_i \in \mathbf{N}(\mathsf{A}) \cap \mathbf{R}(\mathsf{B}).$$

Now, since  $S = {x_1, ..., x_s}$  is a basis for  $N(A) \cap R(B)$  we have

$$\sum_{i=1}^{t} \alpha_i \boldsymbol{z}_i = \sum_{j=1}^{s} \beta_j \boldsymbol{x}_j \quad \Longleftrightarrow \quad$$



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Now, since  $S = \{x_1, \dots, x_s\}$  is a basis for  $N(A) \cap R(B)$  we have

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But  $\mathcal{B} = \{\mathbf{x}_1, \dots, \mathbf{x}_s, \mathbf{z}_1, \dots, \mathbf{z}_t\}$  is linearly independent, so that  $\alpha_1 = \dots = \alpha_t = \beta_1 = \dots = \beta_s = 0$  and



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# Theorem

Let A be an  $m \times n$  matrix, and let B by  $n \times p$ . Then

- rank(AB) ≤ min{rank(A), rank(B)},
- **2**  $rank(AB) \ge rank(A) + rank(B) n.$



We show that  $rank(AB) \le rank(A)$  and  $rank(AB) \le rank(B)$ .



We show that  $rank(AB) \le rank(A)$  and  $rank(AB) \le rank(B)$ .

The previous theorem states

 $rank(AB) = rank(B) - dim(N(A) \cap R(B))$ 



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To make things as tight as possible we take the smaller of the two upper bounds.



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## To prepare for our study of least squares solutions, where the matrices $A^{T}A$ and $AA^{T}$ are important, we prove





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#### Lemma

Let A be a real  $m \times n$  matrix. Then

• rank(
$$A^T A$$
) = rank( $AA^T$ ) = rank( $A$ ).



From our earlier theorem we know

$$\operatorname{rank}(\mathsf{A}^{\mathsf{T}}\mathsf{A}) = \operatorname{rank}(\mathsf{A}) - \operatorname{dim}(\mathsf{N}(\mathsf{A}^{\mathsf{T}}) \cap \mathsf{R}(\mathsf{A})).$$

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## Connection to least squares and normal equations

Consider a — possibly inconsistent — linear system

 $A \boldsymbol{x} = \boldsymbol{b}$ 

with  $m \times n$  matrix A (and  $\boldsymbol{b} \notin R(A)$  if inconsistent).



## Connection to least squares and normal equations

Consider a — possibly inconsistent — linear system

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with  $m \times n$  matrix A (and  $\boldsymbol{b} \notin R(A)$  if inconsistent).

To find a "solution" we multiply both sides by  $A^T$  to get the normal equations:

$$\mathsf{A}^T \mathsf{A} \boldsymbol{x} = \mathsf{A}^T \boldsymbol{b},$$

where  $A^T A$  is an  $n \times n$  matrix.



#### Theorem

Let A be an  $m \times n$  matrix, **b** an m-vector, and consider the normal equations

$$\mathsf{A}^{\mathsf{T}}\mathsf{A}\boldsymbol{x} = \mathsf{A}^{\mathsf{T}}\boldsymbol{b}$$

#### associated with $A \mathbf{x} = \mathbf{b}$ .

- The normal equations are always consistent, i.e., for every A and b there exists at least one x such that A<sup>T</sup>Ax = A<sup>T</sup>b.
- **2** If  $A\mathbf{x} = \mathbf{b}$  is consistent, then  $A^T A\mathbf{x} = A^T \mathbf{b}$  has the same solution set (the least squares solution of  $A\mathbf{x} = \mathbf{b}$ ).
- $A^T A \mathbf{x} = A^T \mathbf{b}$  has a unique solution if and only if rank(A) = n. Then

$$\boldsymbol{x} = (\mathsf{A}^T \mathsf{A})^{-1} \mathsf{A}^T \boldsymbol{b},$$

regardless of whether  $A\mathbf{x} = \mathbf{b}$  is consistent or not.

If  $A\mathbf{x} = \mathbf{b}$  is consistent and has a unique solution, then the same holds for  $A^T A \mathbf{x} = A^T \mathbf{b}$  and  $\mathbf{x} = (A^T A)^{-1} A^T \mathbf{b}$ .

(1) follows from our previous lemma, i.e.,

$$\mathsf{A}^{\mathsf{T}}\boldsymbol{b}\in R(\mathsf{A}^{\mathsf{T}})=R(\mathsf{A}^{\mathsf{T}}\mathsf{A}).$$



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If we multiply by  $A^{T}$ , then

$$\mathsf{A}^T \mathsf{A} \boldsymbol{p} = \mathsf{A}^T \boldsymbol{p},$$

so that **p** is also a solution of the normal equations.

4

### (cont.)

Now, the general solution of  $A\mathbf{x} = \mathbf{b}$  is from the set (see Problem 2 on HW#4)

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$$\boldsymbol{p} + \boldsymbol{N}(\boldsymbol{A}^T\boldsymbol{A}) \stackrel{\text{lemma}}{=} \boldsymbol{p} + \boldsymbol{N}(\boldsymbol{A}) = \mathcal{S}.$$



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Now, if rank( $A^T A$ ) = *n* we know that  $A^T A$  is invertible (even though  $A^T$  and A may not be) and therefore

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$$A^T A \boldsymbol{x} = A^T \boldsymbol{b} \quad \Longleftrightarrow \quad \boldsymbol{x} = (A^T A)^{-1} A^T \boldsymbol{b}.$$

To show (4) we note that  $A\mathbf{x} = \mathbf{b}$  has a unique solution if and only if rank(A) = *n*. But rank(A<sup>T</sup>A) = rank(A) and the rest follows from (3).

#### Remark

The normal equations are not recommended for serious computations since they are often rather ill-conditioned since one can show that

$$\mathit{cond}(\mathsf{A}^{\mathsf{T}}\mathsf{A}) = \mathit{cond}(\mathsf{A})^2.$$

There's an example in [Mey00] that illustrates this fact.



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Example

The matrix

$$\mathsf{A} = \begin{pmatrix} 1 & 2 & 2 & 3 & 1 \\ 2 & 4 & 4 & 6 & 2 \\ 3 & 6 & 6 & 9 & 6 \\ 1 & 2 & 4 & 5 & 3 \end{pmatrix}$$

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cannot have rank 4 since rows one and two are linearly dependent. But rank(A)  $\geq$  2 since  $\begin{pmatrix}9&6\\5&3\end{pmatrix}$  is nonsingular.



$$\label{eq:example} \begin{split} & \text{Example (cont.)} \\ & \text{In fact, rank}(A) = 3 \text{ since} \end{split}$$

$$\begin{pmatrix} 4 & 6 & 2 \\ 6 & 9 & 6 \\ 4 & 5 & 3 \end{pmatrix}$$

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Example (cont.) In fact, rank(A) = 3 since

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is nonsingular.

Note that other singular  $3 \times 3$  submatrices are allowed, such as

$$\begin{pmatrix} 1 & 2 & 2 \\ 2 & 4 & 4 \\ 3 & 6 & 6 \end{pmatrix}$$



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Now

Theorem

Let A and E be  $m \times n$  matrices. Then

 $rank(A + E) \ge rank(A)$ ,

provided the entries of E are "sufficiently small".



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- Beware!! A theoretically singular system may become nonsingular, i.e., have a "solution" — just due to round-off error.
- We may want to intentionally "fix" a singular system, so that it has a "solution". One such strategy is known as Tikhonov regularization, i.e.,

$$\mathsf{A}\mathbf{x} = \mathbf{b} \longrightarrow (\mathsf{A} + \mu \mathsf{I})\mathbf{x} = \mathbf{b},$$

where  $\mu$  is a (small) regularization parameter.



### Proof

We assume that rank(A) = r and that we have nonsingular P and Q such that we can convert A to rank normal form, i.e.,

$$\mathsf{PAQ} = \begin{pmatrix} \mathsf{I}_r & \mathsf{O} \\ \mathsf{O} & \mathsf{O} \end{pmatrix}.$$



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Then — formally — 
$$PEQ = \begin{pmatrix} E_{11} & E_{12} \\ E_{21} & E_{22} \end{pmatrix}$$
 with appropriate blocks  $E_{ij}$ .



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Then — formally —  $PEQ = \begin{pmatrix} E_{11} & E_{12} \\ E_{21} & E_{22} \end{pmatrix}$  with appropriate blocks  $E_{ij}$ . This allows us to write

$$\mathsf{P}(\mathsf{A}+\mathsf{E})\mathsf{Q} = \begin{pmatrix} \mathsf{I}_r + \mathsf{E}_{11} & \mathsf{E}_{12} \\ \mathsf{E}_{21} & \mathsf{E}_{22} \end{pmatrix}.$$



$$(I - B)(I + B + B^2 + ... + B^{k-1})$$



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provided the entries of B are "sufficiently small" (i.e., so that  $B^k \to O$  for  $k \to \infty$ ). Therefore  $(I - B)^{-1}$  exists.

This technique is known as the Neumann series expansion of the inverse of  $\mathsf{I}-\mathsf{B}.$ 



Now, letting  $B = -E_{11}$ , we know that  $(I_r + E_{11})^{-1}$  exists and we can write

$$\begin{pmatrix} I_r & O\\ -E_{21}(I_r + E_{11})^{-1} & I_{m-r} \end{pmatrix} \begin{pmatrix} I_r + E_{11} & E_{12}\\ E_{21} & E_{22} \end{pmatrix} \begin{pmatrix} I_r & -(I_r + E_{11})^{-1}E_{12}\\ O & I_{n-r} \end{pmatrix}$$
$$= \begin{pmatrix} I_r + E_{11} & O\\ O & S \end{pmatrix},$$

where  $S = E_{22} - E_{21}(I_r + E_{11})^{-1}E_{12}$  is the Schur complement of  $I + E_{11}$  in PAQ.



The Schur complement calculation shows that

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But then this rank normal form with invertible diagonal blocks tells us

$$rank(A + E) = rank(I_r + E_{11}) + rank(S)$$



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$$= rank(A) + rank(S)$$



The Schur complement calculation shows that

$$\mathsf{A} + \mathsf{E} \sim \begin{pmatrix} \mathsf{I}_r + \mathsf{E}_{11} & \mathsf{O} \\ \mathsf{O} & \mathsf{S} \end{pmatrix}.$$

But then this rank normal form with invertible diagonal blocks tells us

$$\begin{aligned} \mathsf{rank}(\mathsf{A} + \mathsf{E}) &= \mathsf{rank}(\mathsf{I}_r + \mathsf{E}_{11}) + \mathsf{rank}(\mathsf{S}) \\ &= \mathsf{rank}(\mathsf{A}) + \mathsf{rank}(\mathsf{S}) \\ &\geq \mathsf{rank}(\mathsf{A}). \end{aligned}$$



## Outline



- 2 Four Fundamental Subspaces
- Linear Independence
- 4 Bases and Dimension
- More About Rank
- 6 Classical Least Squares

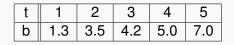


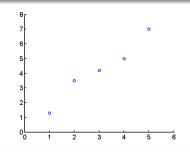


**Classical Least Squares** 

Linear least squares (linear regression)

Given: data  $\{(t_1, b_1), (t_2, b_2), \dots, (t_m, b_m)\}$ 



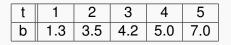


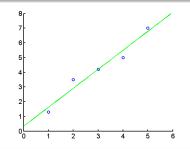


**Classical Least Squares** 

Linear least squares (linear regression)

Given: data  $\{(t_1, b_1), (t_2, b_2), ..., (t_m, b_m)\}$ Find: "best fit" by a line



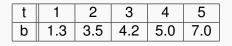


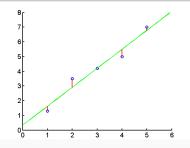


**Classical Least Squares** 

Linear least squares (linear regression)

Given: data  $\{(t_1, b_1), (t_2, b_2), \dots, (t_m, b_m)\}$ Find: "best fit" by a line





### Idea for best fit

Minimize the sum of the squares of the vertical distances of line from the data points.

fasshauer@iit.edu	MATH 532
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$$f(t) = \alpha + \beta t$$



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$$\sum_{i=1}^m \left(f(t_i) - b_i\right)^2$$





$$f(t) = \alpha + \beta t$$

$$\sum_{i=1}^{m} (f(t_i) - b_i)^2$$
$$= \sum_{i=1}^{m} (\alpha + \beta t_i - b_i)^2 \longrightarrow \min$$



$$f(t) = \alpha + \beta t$$

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with  $\alpha$  ,  $\beta$  such that

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From calculus, necessary (and sufficient) condition for minimum

$$rac{\partial {m {G}}(lpha,eta)}{\partial lpha} = {m 0}, \quad rac{\partial {m {G}}(lpha,eta)}{\partial eta} = {m 0}.$$

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where

$$\frac{\partial G(\alpha,\beta)}{\partial \alpha} = 2 \sum_{i=1}^{m} (\alpha + \beta t_i - b_i), \quad \frac{\partial G(\alpha,\beta)}{\partial \beta} = 2 \sum_{i=1}^{m} (\alpha + \beta t_i - b_i) t_i$$

Equivalently,

$$\left(\sum_{i=1}^{m} 1\right) \alpha + \left(\sum_{i=1}^{m} t_i\right) \beta = \sum_{i=1}^{m} b_i$$
$$\left(\sum_{i=1}^{m} t_i\right) \alpha + \left(\sum_{i=1}^{m} t_i^2\right) \beta = \sum_{i=1}^{m} b_i t_i$$



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with

$$\mathbf{Q} = \begin{pmatrix} \sum_{i=1}^{m} 1 & \sum_{i=1}^{m} t_i \\ \sum_{i=1}^{m} t_i & \sum_{i=1}^{m} t_i^2 \end{pmatrix}, \quad \mathbf{x} = \begin{pmatrix} \alpha \\ \beta \end{pmatrix}, \quad \mathbf{y} = \begin{pmatrix} \sum_{i=1}^{m} b_i \\ \sum_{i=1}^{m} b_i t_i \end{pmatrix}$$



**Classical Least Squares** 

We can write each of these sums as inner products:



$$\sum_{i=1}^{m} \mathbf{1} = \mathbf{1}^{T} \mathbf{1}, \quad \sum_{i=1}^{m} t_{i} = \mathbf{1}^{T} t = t^{T} \mathbf{1}, \quad \sum_{i=1}^{m} t_{i}^{2} = t^{T} t$$
$$\sum_{i=1}^{m} b_{i} = \mathbf{1}^{T} b = b^{T} \mathbf{1}, \quad \sum_{i=1}^{m} b_{i} t_{i} = b^{T} t = t^{T} b,$$

where

$$\mathbf{1}^{T} = \begin{pmatrix} 1 & \cdots & 1 \end{pmatrix}, \quad \mathbf{t}^{T} = \begin{pmatrix} t_{1} & \cdots & t_{m} \end{pmatrix}, \quad \mathbf{b}^{T} = \begin{pmatrix} b_{1} & \cdots & b_{m} \end{pmatrix}$$



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With this notation we have

$$Q \boldsymbol{x} = \boldsymbol{y} \iff$$



$$\sum_{i=1}^{m} \mathbf{1} = \mathbf{1}^{T} \mathbf{1}, \quad \sum_{i=1}^{m} t_{i} = \mathbf{1}^{T} t = t^{T} \mathbf{1}, \quad \sum_{i=1}^{m} t_{i}^{2} = t^{T} t$$
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With this notation we have

$$\mathbf{Q}\mathbf{x} = \mathbf{y} \quad \Longleftrightarrow \quad \begin{pmatrix} \mathbf{1}^T \mathbf{1} & \mathbf{1}^T \mathbf{t} \\ \mathbf{t}^T \mathbf{1} & \mathbf{t}^T \mathbf{t} \end{pmatrix} \mathbf{x} = \begin{pmatrix} \mathbf{1}^T \mathbf{b} \\ \mathbf{t}^T \mathbf{b} \end{pmatrix}$$



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$$\iff A^{T}A\mathbf{x} = A^{T}\mathbf{b}, \qquad A^{T} = \begin{pmatrix} \mathbf{1}^{T} \\ \mathbf{t}^{T} \end{pmatrix}, \quad A = (\mathbf{1} \quad \mathbf{t})$$

$$A^T A \boldsymbol{x} = A^T \boldsymbol{b}.$$



$$A^T A \boldsymbol{x} = A^T \boldsymbol{b}.$$

Also note that since  $\varepsilon_i = \alpha + \beta t_i - b_i$  we have

$$\boldsymbol{\varepsilon} = \begin{pmatrix} \varepsilon_1 \\ \vdots \\ \varepsilon_m \end{pmatrix} = \begin{pmatrix} 1 \\ \vdots \\ 1 \end{pmatrix} \alpha + \begin{pmatrix} t_1 \\ \vdots \\ t_m \end{pmatrix} \beta - \begin{pmatrix} b_1 \\ \vdots \\ b_m \end{pmatrix}$$
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This implies that

$$G(\alpha,\beta) = \sum_{i=1}^{m} \varepsilon_i^2$$



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$$= \mathbf{1}\boldsymbol{\alpha} + \mathbf{t}\boldsymbol{\beta} - \mathbf{b} = \mathbf{A}\mathbf{x} - \mathbf{b}.$$

This implies that

$$G(\alpha, \beta) = \sum_{i=1}^{m} \varepsilon_i^2 = \varepsilon^T \varepsilon = (\mathbf{A} \mathbf{x} - \mathbf{b})^T (\mathbf{A} \mathbf{x} - \mathbf{b})^T$$



ſ	t	-1	0	1	2	3	4	5	6
	b	10	9	7	5	4	3	0	-1

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 $A^T A \boldsymbol{x} = A^T \boldsymbol{b}$ 

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$$A^{T}A\boldsymbol{x} = A^{T}\boldsymbol{b} \iff \begin{pmatrix} \sum_{i=1}^{8} 1 & \sum_{i=1}^{8} t_{i} \\ \sum_{i=1}^{8} t_{i} & \sum_{i=1}^{8} t_{i}^{2} \end{pmatrix} \begin{pmatrix} \alpha \\ \beta \end{pmatrix} = \begin{pmatrix} \sum_{i=1}^{8} b_{i} \\ \sum_{i=1}^{8} b_{i} t_{i} \end{pmatrix}$$
$$\iff \begin{pmatrix} 8 & 20 \\ 20 & 92 \end{pmatrix} \begin{pmatrix} \alpha \\ \beta \end{pmatrix} = \begin{pmatrix} 37 \\ 25 \end{pmatrix}$$
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Data:

$$A^{T}A\mathbf{x} = A^{T}\mathbf{b} \quad \Longleftrightarrow \quad \begin{pmatrix} \sum_{i=1}^{8} 1 & \sum_{i=1}^{8} t_{i} \\ \sum_{i=1}^{8} t_{i} & \sum_{i=1}^{8} t_{i}^{2} \end{pmatrix} \begin{pmatrix} \alpha \\ \beta \end{pmatrix} = \begin{pmatrix} \sum_{i=1}^{8} b_{i} \\ \sum_{i=1}^{8} b_{i} t_{i} \end{pmatrix}$$
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So that the best fit line to the given data is

$$f(t) \approx 8.643 - 1.607t.$$

# **General Least Squares**

The general least squares problem behaves analogously to the linear example.

#### Theorem

Let A be a real  $m \times n$  matrix and **b** an m-vector. Any vector **x** that minimizes the square of the residual A $\mathbf{x} - \mathbf{b}$ , i.e.,

$$G(\boldsymbol{x}) = (A\boldsymbol{x} - \boldsymbol{b})^T (A\boldsymbol{x} - \boldsymbol{b})$$

is called a least squares solution of Ax = b.

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is called a least squares solution of Ax = b.

The set of all least squares solutions is obtained by solving the normal equations

$$\mathsf{A}^{\mathsf{T}}\mathsf{A}\boldsymbol{x} = \mathsf{A}^{\mathsf{T}}\boldsymbol{b}.$$

Moreover, a unique solution exists if and only if rank(A) = n so that

$$\boldsymbol{x} = (\mathsf{A}^T \mathsf{A})^{-1} \mathsf{A}^T \boldsymbol{b}.$$

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=  $\mathbf{x}^{\mathsf{T}} \mathbf{A}^{\mathsf{T}} \mathbf{A}\mathbf{x} - 2\mathbf{x}^{\mathsf{T}} \mathbf{A}^{\mathsf{T}} \mathbf{b} + \mathbf{b}^{\mathsf{T}} \mathbf{b}$ 

## Proof

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since  $\boldsymbol{b}^T A \boldsymbol{x} = (\boldsymbol{b}^T A \boldsymbol{x})^T = \boldsymbol{x}^T A^T \boldsymbol{b}$  is a scalar.

$$\frac{\partial G(\boldsymbol{x})}{\partial x_i} = \frac{\partial \boldsymbol{x}^T}{\partial x_i} \mathsf{A}^T \mathsf{A} \boldsymbol{x} + \boldsymbol{x}^T \mathsf{A}^T \mathsf{A} \frac{\partial \boldsymbol{x}}{\partial x_i} - 2 \frac{\partial \boldsymbol{x}^T}{\partial x_i} \mathsf{A}^T \boldsymbol{b}$$

$$\frac{\partial G(\boldsymbol{x})}{\partial x_i} = \frac{\partial \boldsymbol{x}^T}{\partial x_i} A^T A \boldsymbol{x} + \boldsymbol{x}^T A^T A \frac{\partial \boldsymbol{x}}{\partial x_i} - 2 \frac{\partial \boldsymbol{x}^T}{\partial x_i} A^T \boldsymbol{b}$$
$$= \boldsymbol{e}_i^T A^T A \boldsymbol{x} + \boldsymbol{x}^T A^T A \boldsymbol{e}_i - 2 \boldsymbol{e}_i^T A^T \boldsymbol{b}$$

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$$= \boldsymbol{e}_i^T A^T A \boldsymbol{x} + \boldsymbol{x}^T A^T A \boldsymbol{e}_i - 2 \boldsymbol{e}_i^T A^T \boldsymbol{b}$$
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since  $\mathbf{x}^T A^T A \mathbf{e}_i = (\mathbf{x}^T A^T A \mathbf{e}_i)^T = \mathbf{e}_i^T A^T A \mathbf{x}$  is a scalar. This means that

$$\frac{\partial G(\boldsymbol{x})}{\partial x_i} = 0 \quad \Longleftrightarrow \quad$$

$$\frac{\partial G(\boldsymbol{x})}{\partial x_i} = \frac{\partial \boldsymbol{x}^T}{\partial x_i} A^T A \boldsymbol{x} + \boldsymbol{x}^T A^T A \frac{\partial \boldsymbol{x}}{\partial x_i} - 2 \frac{\partial \boldsymbol{x}^T}{\partial x_i} A^T \boldsymbol{b}$$
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If we collect all such conditions (for i = 1, ..., n) in one linear system we get

$$\mathsf{A}^{\mathsf{T}}\mathsf{A}\boldsymbol{x} = \mathsf{A}^{\mathsf{T}}\boldsymbol{b}.$$

$$G(\boldsymbol{z}) = (\boldsymbol{A}\boldsymbol{z} - \boldsymbol{b})^T (\boldsymbol{A}\boldsymbol{z} - \boldsymbol{b}) \\ = \boldsymbol{z}^T \boldsymbol{A}^T \boldsymbol{A}\boldsymbol{z} - 2\boldsymbol{z}^T \boldsymbol{A}^T \boldsymbol{b} + \boldsymbol{b}^T \boldsymbol{b}$$

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To verify that we indeed have a minimum we show that if z is a solution of the normal equations then G(z) is minimal.

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=  $z^{T}(\underbrace{A^{T}Az - A^{T}b}_{=0}) - z^{T}A^{T}b + b^{T}b = -z^{T}A^{T}b + b^{T}b.$   
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$$G(\mathbf{y}) = (\mathbf{z} + \mathbf{u})^T \mathsf{A}^T \mathsf{A}(\mathbf{z} + \mathbf{u}) - 2(\mathbf{z} + \mathbf{u})^T \mathsf{A}^T \mathbf{b} + \mathbf{b}^T \mathbf{b}$$

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Now, for any other y = z + u we have

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since  $\boldsymbol{u}^T A^T A \boldsymbol{u} = \sum_{i=1}^m (A \boldsymbol{u})_i^2 \ge 0$ .  $\Box$ 

Using this framework we can compute least squares fits from any linear function space.



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#### Example

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- Solution Let  $f(t) = \alpha e^t + \beta \sqrt{t}$ , i.e., we can use just about anything we want.



# Regression in Statistics (BLUE)

One assumes that there is a random process that generates data as a random variable Y of the form

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \ldots + \beta_n X_n,$$

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Now the actually observed data may be affected by noise, i.e.,

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$$y_i = \beta_0 + \beta_1 x_{i,1} + \beta_2 x_{i,2} + \ldots + \beta_n x_{i,n} + \varepsilon, \quad i = 1, \ldots, m.$$



$$\boldsymbol{y} = \boldsymbol{X}\boldsymbol{eta} + \boldsymbol{arepsilon}$$

Now, the least squares solution of  $X\beta = y$ , i.e.,  $\hat{\beta} = (X^T X)^{-1} X^T y$  is in fact the best linear unbiased estimator (BLUE) for  $\beta$ .



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so that the estimator is indeed unbiased.



One can also show (maybe later) that  $\hat{\beta}$  has minimal variance among all unbiased linear estimators, so it is the best linear unbiased estimator of the model parameters.

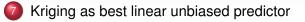
In fact, the theorem ensuring this is the so-called Gauss-Markov theorem.



# Outline



- 2) Four Fundamental Subspaces
- Linear Independence
- 4 Bases and Dimension
- 5 More About Rank
- 6 Classical Least Squares





**Assume:** the approximate value of a realization of a zero-mean (Gaussian) random field is given by a linear predictor of the form

$$\hat{Y}_{\boldsymbol{x}} = \sum_{j=1}^{N} Y_{\boldsymbol{x}_{j}} w_{j}(\boldsymbol{x}) = \boldsymbol{w}(\boldsymbol{x})^{T} \boldsymbol{Y},$$

where  $\hat{Y}_{\boldsymbol{x}}$  and  $Y_{\boldsymbol{x}_j}$  are random variables,  $\boldsymbol{Y} = (Y_{\boldsymbol{x}_1} \cdots Y_{\boldsymbol{x}_N})^T$ , and  $\boldsymbol{w}(\boldsymbol{x}) = (w_1(\boldsymbol{x}) \cdots w_N(\boldsymbol{x}))^T$  is a vector of weight functions at  $\boldsymbol{x}$ .



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$$\mathsf{MSE}(\hat{Y}_{\boldsymbol{x}}) = \mathbb{E}\left[\left(Y_{\boldsymbol{x}} - \boldsymbol{w}(\boldsymbol{x})^{\mathsf{T}}\boldsymbol{Y}\right)^{2}\right].$$



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We now present some details (see [FM15]).



$$\sigma^{2} \mathcal{K}(\boldsymbol{x}, \boldsymbol{z}) = \operatorname{Cov}(Y_{\boldsymbol{x}}, Y_{\boldsymbol{z}}) = \mathbb{E}\left[(Y_{\boldsymbol{x}} - \mu(\boldsymbol{x}))(Y_{\boldsymbol{z}} - \mu(\boldsymbol{z}))\right]$$



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=  $\mathbb{E}\left[(Y_{\mathbf{x}} - \mathbb{E}[Y_{\mathbf{x}}])(Y_{\mathbf{z}} - \mathbb{E}[Y_{\mathbf{z}}])\right]$   
=  $\mathbb{E}\left[Y_{\mathbf{x}}Y_{\mathbf{z}} - Y_{\mathbf{x}}\mathbb{E}[Y_{\mathbf{z}}] - \mathbb{E}[Y_{\mathbf{x}}]Y_{\mathbf{z}} + \mathbb{E}[Y_{\mathbf{x}}]\mathbb{E}[Y_{\mathbf{z}}]\right]$ 



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$$= \mathbb{E}\left[Y_{\mathbf{x}}Y_{\mathbf{z}} - Y_{\mathbf{x}}\mathbb{E}[Y_{\mathbf{z}}] - \mathbb{E}[Y_{\mathbf{x}}]Y_{\mathbf{z}} + \mathbb{E}[Y_{\mathbf{x}}]\mathbb{E}[Y_{\mathbf{z}}]\right]$$
  
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We need the covariance kernel *K* of a random field *Y* with mean  $\mu(\mathbf{x})$ . It is defined via

$$\sigma^{2} \mathcal{K}(\mathbf{x}, \mathbf{z}) = \operatorname{Cov}(Y_{\mathbf{x}}, Y_{\mathbf{z}}) = \mathbb{E}\left[(Y_{\mathbf{x}} - \mu(\mathbf{x}))(Y_{\mathbf{z}} - \mu(\mathbf{z}))\right]$$
  
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Therefore, the variance of the random field,

$$\operatorname{Var}(Y_{\boldsymbol{X}}) = \mathbb{E}[Y_{\boldsymbol{X}}^2] - \mathbb{E}[Y_{\boldsymbol{X}}]^2 = \mathbb{E}[Y_{\boldsymbol{X}}^2] - \mu^2(\boldsymbol{X}),$$

corresponds to the "diagonal" of the covariance, i.e.,

$$\operatorname{Var}(Y_{\boldsymbol{x}}) = \sigma^2 K(\boldsymbol{x}, \boldsymbol{x}).$$



$$\mathsf{MSE}(\hat{Y}_{\boldsymbol{x}}) = \mathbb{E}\left[\left(Y_{\boldsymbol{x}} - \boldsymbol{w}(\boldsymbol{x})^{T} \boldsymbol{Y}\right)^{2}\right]$$



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Now use  $\mathbb{E}[Y_x Y_z] = K(x, z)$  (the covariance, since Y is centered):

$$\mathsf{MSE}(\hat{Y}_{\boldsymbol{x}}) = \sigma^2 \mathcal{K}(\boldsymbol{x}, \boldsymbol{x}) - 2\boldsymbol{w}(\boldsymbol{x})^T (\sigma^2 \boldsymbol{k}(\boldsymbol{x})) + \boldsymbol{w}(\boldsymbol{x})^T (\sigma^2 \mathsf{K}) \boldsymbol{w}(\boldsymbol{x}),$$

where

$$\sigma^{2}\boldsymbol{k}(\boldsymbol{x}) = \sigma^{2} \begin{pmatrix} k_{1}(\boldsymbol{x}) & \cdots & k_{N}(\boldsymbol{x}) \end{pmatrix}^{T} : \text{ with} \\ \sigma^{2}k_{j}(\boldsymbol{x}) = \sigma^{2}K(\boldsymbol{x}, \boldsymbol{x}_{j}) = \mathbb{E}[Y_{\boldsymbol{x}}Y_{\boldsymbol{x}_{j}}] \\ \text{K: the covariance matrix has entries } \sigma^{2}K(\boldsymbol{x}_{i}, \boldsymbol{x}_{j}) = \mathbb{E}[Y_{\boldsymbol{x}_{i}}Y_{\boldsymbol{x}_{i}}]$$



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K: the covariance matrix has entries  $\sigma^2 K(\mathbf{x}_i, \mathbf{x}_j) = \mathbb{E}[Y_{\mathbf{x}_i} Y_{\mathbf{x}_j}]$ Finding the minimum MSE is straightforward.



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$$-2\boldsymbol{k}(\boldsymbol{x})+2\mathsf{K}\boldsymbol{w}(\boldsymbol{x})=0,$$

and so the optimum weight vector is

$$\overset{\star}{\boldsymbol{w}}(\boldsymbol{x}) = \mathsf{K}^{-1}\boldsymbol{k}(\boldsymbol{x}).$$



We have shown that the (simple) kriging predictor

$$\hat{Y}_{\boldsymbol{x}} = \boldsymbol{k}(\boldsymbol{x})^T \mathsf{K}^{-1} \boldsymbol{Y}$$

is the best (in the MSE sense) linear unbiased predictor (BLUP).



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is the best (in the MSE sense) linear unbiased predictor (BLUP).

Since we are given the observations y as realizations of Y we can compute the prediction

$$\hat{y}_{\boldsymbol{x}} = \boldsymbol{k}(\boldsymbol{x})^T \mathbf{K}^{-1} \boldsymbol{y}.$$



The MSE of the kriging predictor with optimal weights  $\dot{w}(\cdot)$ ,

$$\mathbb{E}\left[\left(\boldsymbol{Y}_{\boldsymbol{x}}-\hat{\boldsymbol{Y}}_{\boldsymbol{x}}\right)^{2}\right]=\sigma^{2}\left(\boldsymbol{K}(\boldsymbol{x},\boldsymbol{x})-\boldsymbol{k}(\boldsymbol{x})^{T}\boldsymbol{K}^{-1}\boldsymbol{k}(\boldsymbol{x})\right),$$

is known as the kriging variance.

It allows us to give confidence intervals for our prediction. It also gives rise to a criterion for choosing an optimal parametrization of the family of covariance kernels used for prediction.



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#### Remark

For Gaussian random fields the BLUP is also the best nonlinear unbiased predictor (see, e.g., [BTA04, Chapter 2]).



#### Remark

- The simple kriging approach just described is precisely how Krige [Kri51] introduced the method:
  - The unknown value to be predicted is given by a weighted average of the observed values, where the weights depend on the prediction location.
  - Usually one assigns a smaller weight to observations further away from **x**.



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The latter statement implies that one should be using kernels whose associated weights decay away from  $\mathbf{x}$ . Positive definite translation invariant kernels have this property.



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More advanced kriging variants are discussed in papers such as [SWMW89, SSS13], or books such as [Cre93, Ste99, BTA04].



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